


## Article

# Do Urbanization and Energy Consumption Change the Role in Environmental Degradation in the European Union Countries?

Bartosz Jóźwik , Antonina-Victoria Gavryshkiv and Kinga Galewska

Department of International Economics, The John Paul II Catholic University of Lublin, Al. Raclawickie 14, 20-950 Lublin, Poland

\* Correspondence: bjozwik@kul.pl

**Abstract:** Nowadays, increased urbanization is visible in most European Union countries. At the same time, it can be noticed that in the studied period (2000–2018), GDP per capita increased, and CO<sub>2</sub> emissions per capita and energy consumption per capita decreased. These trends should be assessed in an unequivocally positive way. Considering these trends, especially with regard to economic development, our research goal is to answer the following questions: is there a long-run relationship between urbanization, energy consumption, economic growth, and carbon dioxide emissions, and what roles do urbanization and energy consumption play in the concept of the environmental Kuznets curve? This study aims to contribute to this growing area of research by exploring the European Union countries in the period covering the accession of new member states from Central Europe that needs intensifying European environmental policy. In order to test cointegration, we used Pedroni and Westerlund's panel tests. To estimate the long-run coefficients, we employed the FMOLS, MG, CCEMG, and AMG tests. Our findings confirmed the long-run relationship between variables. We find that urbanization has a high negative impact on carbon dioxide emissions per capita. Interestingly, our studies' results differ from those in most of the previously published articles about European countries. For this reason, our results provide a new insight for policymakers in European Union institutions.

**Keywords:** environmental Kuznets curve; carbon dioxide emissions; CO<sub>2</sub>; urbanization; energy consumption; European Union



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## 1. Introduction

Nowadays, over half of the world's inhabitants live in cities. The United Nations' prognoses point out that the total population in the world in 2050 will reach 9.31 billion, while the urban population will increase to 6.25 billion, and the urbanization index will be 67.2% [1]. It is largely the civilization advance, together with all the accompanying effects, which has made the population in cities grow dramatically. However, all these aspects and assumptions have consequences as far as the natural environment, the population growth, and the population distribution in particular areas of the globe, especially in cities, are concerned. In the European Union countries, urbanization is progressing continuously, extending into new regions. In the years 2000–2018, its increase was visible in most countries. The urbanization index dropped only in four countries: Slovak Republic, Austria, Cyprus, and Poland. Minor changes (less than 1%) in this respect occurred only in the Baltic countries (Estonia, Latvia, and Lithuania), the Czech Republic, and Belgium (Table A1 in Appendix A). Based on World Bank Statista data, in 2019, 75% of the population lived in cities and the suburbs of the European Union countries, while only 25% lived in rural areas. It is noteworthy that at the same time, a decrease in carbon dioxide emissions can be observed along with the process of growing urbanization.

In most countries, a decreasing carbon dioxide emissions tendency is seen in comparing the emissions in the years 2000 and 2018. For example, Luxembourg, the leader

in this ranking in 2000, witnessed significant changes in the field of applied policy concerning emissions reduction. In this region, the estimated decrease in emissions was from 19.7 metric tons per capita in 2000 to 15.3 in 2018 (Table A1 in Appendix A). The visible changes (more than 1%), in this respect, proceed in another direction in the Baltic countries (Lithuania, Latvia, and Estonia). There, an increase in carbon dioxide emissions took place by 1.03, 1.13, and 1.50 metric tons per capita, respectively. However, it must be noted that in most countries, the ongoing changes are an element of the idea of climate neutrality, which is the aim of the European Union for the next decades.

The European Union treats the problem of climate change in a very emphatic way, and it undertakes activities in this direction. Prevention of those changes is one of its priority goals, reflected in the tasks designed for the decades to come, for example, through a reduction in greenhouse gas emissions [2]. The European Green Deal is a strategy for growth, transforming the economic and political union of 27 European democratic countries into places that are neutral to climate. The activities accompanying the major goal refer to significant aspects. Firstly, it is the establishment of a modern, resource-efficient, and competitive economy where there will be no net emissions of greenhouse gases in 2050. Secondly, it is a separation of economic growth from the use of resources. The third aspect refers to guaranteeing the protection and strengthening of neutral capital. Finally, the fourth, but nonetheless very important point, is to ensure citizens' health protection, security, and well-being, which is aimed at protecting them from the environmental effects of climate change.

Considering the formulated aims of climate neutrality, as well as the economic development and the progressing process of urbanization in the European Union countries, our main research goal is to answer the question, is there a long-run relationship between urbanization, energy consumption, economic growth, and the carbon dioxide emissions, and what roles do urbanization and energy consumption play in the concept of the environmental Kuznets curve? This study aims to contribute to this growing area of research by exploring the European Union countries in the period, which covers the accession of new member states from Central Europe. This enlargement needs intensifying cooperation between EU member states, especially in environmental policy. The relationship is tested using the concept of the environmental Kuznets curve, where, apart from carbon dioxide emissions and economic growth, urbanization and final energy consumption are considered. To this aim, the Pedroni and Westerlund panel cointegration tests are used. To estimate the long-run coefficients of the cointegration association, we employed the panel Fully Modified Ordinary Least Squares (FMOLS) test. To test the robustness of the estimation results, we used the Pesaran and Smith Mean Group (MG) estimator, the Pesaran Common Correlated Effects Mean Group (CCEMG), and Augmented Mean Group (AMG) estimators.

The remaining sections of this research are planned as follows. Section 2 presents a brief literature review on the relationship between urbanization, environmental degradation, and economic growth, analyzed within the environmental Kuznets curve (EKC) concept. Section 3 contains the data, model and empirical methodology. The research results and discussion are presented in Section 4. Section 5 concludes the research and provides policy recommendations.

## 2. Literature Review

Research on the effect of urbanization on the quality of the environment is frequently conducted using the concept of the environmental Kuznets curve, which appeared at the beginning of the 1990s in the work by Grossman and Krueger [3]. The Authors proved that the scale of environmental pollution is connected with the level of economic development of a given country. In the initial stages of economic development, an increase in the level of pollution related to the exploitation of natural resources also takes place, intending to create welfare. This tendency is reversed after a certain level of income (turning point) is trespassed. Then, the situation changes, and expenditures on environmental protection start to increase. The conclusions drawn by Grossman and Krueger became the basis for

creating a model according to which the relationships between economic growth and the emissions of pollutants have an inverted U-shaped curve. In recent years, the popularity of the environmental Kuznets curve, which additionally used different variables, grew. A complex review of the literature in this area can be found, for instance, in Shahbaz and Sinha [4,5], Purcel [6], Koondhar et al. [7], and Xia et al. [8]. This article focuses on research that primarily considers the variables characterizing the urbanization process, urban population, and energy consumption.

It needs to be emphasized that this research was carried out in various regions and states with different levels of economic development, for instance, in emerging economies, developing countries, or developed countries. Most of those studies confirm the relationships defined by the environmental Kuznets curve, but the results of the effect of urbanization on the quality of the environment are not conclusive. For example, Maneejuk et al. [9] analyzed the relationship between GDP per capita, urbanization, financial development, the industrial sector, and the emissions of CO<sub>2</sub> for the Association of Southeast Asian Nations (ASEAN), the European Free Trade Association (EFTA), the European Union (EU), Group of Seven (G7), Gulf Cooperation Council (GCC), Mercosur, the North American Free Trade Agreement (NAFTA) and the Organization for Economic Co-operation and Development (OECD) in the years 2001–2016. The findings indicate that the EKC hypothesis is valid in only three out of eight international economic communities, namely, the EU, OECD, and G7. It follows from the research that urbanization, as well as financial development and the industrial sector, increase CO<sub>2</sub> emissions, while the use of renewable energy reduces degradation of the environment. In the case of urbanization, statistical significance and the highest positive effect were displayed by ASEAN (0.823), and then by GCC (0.563), Mercosur (0.553), UEU (0.123), and G7 (0.019). In the other groups, the effect of urbanization on CO<sub>2</sub> emissions was statistically insignificant.

Similar results were obtained by Wang et al. [1]. The Authors analyzed the effect of urbanization on economic growth and the quality of the environment in the period 1996–2015 based on data from 134 countries. Studies confirmed the occurrence of an inverted U-shaped relationship between economic growth and CO<sub>2</sub> emissions for the countries in the lower middle-income group, and a U-shaped relationship for the high-income group of countries. The Authors showed that the emissions of CO<sub>2</sub> increased together with increased urbanization. The same direction of the effect of urbanization on carbon dioxide emissions was defined by Sun Y. et al. [10], who conducted research on the Middle East and North African (MENA) economies.

However, it deserves to be pointed out that an increase in urbanization can increase the emissions of carbon dioxide only to a certain level, after which its further progress will reduce these emissions. Such relationships were confirmed in the studies by Gierałtowska et al. [11], who indicated that urbanization has an inverted U-shaped relationship with CO<sub>2</sub> emissions in the group covering 163 countries over the period from 2000 to 2016. This relationship can be confirmed by the results of studies obtained by Li and Haneklaus [12], where the Authors showed that increased urbanization decreases CO<sub>2</sub> emissions in the group of G7 countries in the years 1979 to 2019, and by Balsalobre-Lorente et al. [13] in the BRICS states in the years 1914–2014. Likewise, studies by Saidi and Mbarek [14], on the effect of urbanization, income, trade openness, and financial development on the carbon dioxide emissions in nineteen emerging economies during 1990–2013, indicate that urbanization decreases CO<sub>2</sub> emissions. According to the Authors, this is a powerful argument for politicians and city planners in shaping contemporary policies in those regions.

Previous studies conducted in European countries indicated the opposite results. Based on research carried out in 33 European countries and covering the period 1996–2017, Ali et al. [15] showed that urbanization together with economic growth, export, import, and energy consumption are the main factors that increase environmental degradation. The coefficients associated with urbanization are positive and statistically significant: 0.188 for model I, and 0.011 for model II. At the same time, the Authors point to energy innovation,

which should help to reduce the rate of environmental degradation. A similar group of European countries (36) was studied by Wang et al. [16], who indicated a positive and significant effect of urbanization, as well as economic growth and foreign direct investment, on CO<sub>2</sub> emissions in the years 2000–2018. A slightly bigger group was examined by Khezri et al. [17]. The results of studies for 43 European countries between 1996 and 2018 also confirmed the relationship defined as the environmental Kuznets curve and urbanization's positive effect on carbon dioxide emissions (coefficients 0.659–0.760).

Comparable results, but for smaller groups of European countries, were obtained by Balsalobre-Lorente et al. [18]. The Authors studied the relationships between GDP per capita, urbanization, foreign direct investment, renewable energy consumption, and CO<sub>2</sub> emissions in Portugal, Ireland, Italy, Greece, and Spain, in the years 1990–2019. The study confirmed the relationship between economic growth and CO<sub>2</sub> emissions in the inverted U-shaped and N-shaped curves. The urbanization process increases the emissions of CO<sub>2</sub> in such a way that an increase in urbanization by 1% increases CO<sub>2</sub> emissions within the range from 0.44% to 6.36%, depending on the adopted model. Verbič et al. [19] conducted studies for the countries of South-Eastern Europe in the years 1997–2014. The evidence points to an inverted U-shaped relationship between GDP per capita and the emissions of carbon dioxide in the long run in the whole sample. Short-term estimates evidence the existence of EKC in the inverted U-shape only for Greece and Moldavia. The Authors pointed to a statistically significant positive influence of urbanization on CO<sub>2</sub> emissions (coefficient 1.057, FMOLS).

However, not all studies confirm the negative or positive effects of urbanization on the emissions of carbon dioxide. To give an example, no relation between urbanization and carbon dioxide emissions was indicated by Destek et al. [20]. Their research sample comprised Central European countries such as Albania, Bulgaria, Croatia, the Czech Republic, Macedonia, Hungary, Poland, Romania, Slovakia, and Slovenia. The main goal was to find the relationship between CO<sub>2</sub> emissions, urbanization, GDP per capita, energy consumption, and trade openness in the years 1991–2011. Studies confirmed the hypothesis of the environmental Kuznets curve in the sample. Results indicate a short-run two-directional causal relation between CO<sub>2</sub> and GDP per capita as well as between GDP per capita and energy consumption. There is no relation, however, between urbanization and carbon dioxide emissions. Similar results were obtained by Amin, et al. [21], who point out that urbanization in European countries does have a positive effect on environmental pollution, but it is statistically insignificant. Interestingly, the Authors saw a need to analyze the transport sector as a consequence of the process of urbanization. The Authors argue that transport significantly affects the air quality. They also point out that using renewable energy reduces carbon dioxide emissions from transportation. At the same time, they emphasize that necessary measures should be taken to increase ecological consciousness, especially among the urban population. In this process, it is important to promote environmentally friendly and energy-efficient means of transportation.

Although, the impact of urbanization on the environment in the European Union is related to the fact that some countries have undergone deindustrialization and offshored the environmental effects of their consumption to other countries. Research on industrialization's impact on carbon dioxide emissions mainly focuses on structural changes, where structural changes towards services, usually at higher levels of economic development, improve environmental quality [22–24]. Previous works, including Cherniwchan [25], and Raheem and Ogebe [26], have shown that industrialization is an important determinant in environmental quality changes. Another problem is offshoring the negative ecological impacts, which is often the result of differences in carbon prices in different regions. This phenomenon can lead to the production of energy-intensive goods into “carbon havens”, thus creating a “carbon leakage”. The observed industry relocation is a significant problem for the European Union and national policymakers [27].

### 3. Materials and Methods

We use the model that characterizes the relationships between economic development and the degree of environmental pollution. The first studies on these relationships included those by Grossman and Krueger [3,28], Shafik and Bandyopadhyay [29], Panayotou [30], and Selden and Song [31]. A fast increase in the number of studies led to the formulation of the concept of the environmental Kuznets curve, for example, see Gruszecki and Jóźwik [32]. It assumes a relationship between economic escalation (GDP per capita) and the level of nature contamination (e.g., due to carbon dioxide emissions), mostly in the inverted U-shaped curve. It happens because industrialization is followed by certain negative consequences (for example, pollution of man's natural environment), which grow to a certain point, after which they decrease, even though economic development proceeds. This, on the other hand, follows on from the fact that at a certain stage of advanced economic development, a change can be noticed in the mechanism of demands exhibited by consumers who then, to function, need more services and a cleaner environment. Technological progress also takes place, which does away with the negative effects of contamination of the surrounding world following economic development.

Considering the realization of our research goal, an important problem proves to be the aforementioned relationship described by the environmental Kuznets curve and the observed increase in urbanization and technological progress in the European Union countries. This relationship induces a search for the answer to the following question: is there a long-run relationship between urbanization, energy consumption, economic growth, and carbon dioxide emissions, and what roles do urbanization and energy consumption play in the concept of the environmental Kuznets curve? We use the econometric model with the urbanization variable to answer the questions. The model with the urbanization variable was employed, for example, by Kasman and Duman [33], Ozatac, Gokmenoglu, and Taspinar [34] as well as by Kirikkaleli and Kalmaz [35], Musa et al. [36], and Anwar et al. [37]. Our model will also consider final energy consumption.

The relationship between carbon dioxide emissions, GDP per capita, urban population (urbanization), and final energy consumption per capita is expressed in model I (Equation (1)). We also use model II (Equation (2)) for a robustness check where environmental degradation is proxied as greenhouse gas emissions, expressed in units of CO<sub>2</sub> equivalents. All variables are transformed into a natural logarithm format, to avoid multicollinearity issues, reduce the possible outliers from the dataset, as well as overcome the chances of data sharpness and normality [13].

$$\ln\text{CO}_{2it} = \beta_0 + \beta_1 \ln \text{GDP} + \beta_2 (\ln \text{GDP})^2 + \beta_3 \ln \text{URB} + \beta_4 \ln \text{ENC} + \mu_{it} \quad (1)$$

$$\ln\text{GHG}_{it} = \beta_0 + \beta_1 \ln \text{GDP} + \beta_2 (\ln \text{GDP})^2 + \beta_3 \ln \text{URB} + \beta_4 \ln \text{ENC} + \mu_{it} \quad (2)$$

where  $\beta$ —regression coefficients, CO<sub>2</sub>—carbon dioxide emissions in metric tons per capita, GHG—greenhouse gas emissions per capita, GDP—gross domestic product per capita (constant 2015 USD), URB—urban population (% of total population), ENC—final energy consumption in tonnes of oil equivalent per capita,  $\mu_{it}$ —error correction term. It should be pointed out that in many scientific studies, the standard measure of urbanization is the share of the population living in urban areas [38].

Before estimating the models, some preliminary tests need to be applied to the panel data. Figure 1 shows the entire research procedure. Initially, we test for cross-section dependence using the Pesaran CD-test [39]. Afterward, in order to discover whether the data of selected variables have stationarity or non-stationarity, we apply the Im–Pesaran–Shin panel unit root test [40] and the second-generation unit root test in the presence of cross-section dependence proposed by Pesaran [41]. Next, we test the long-run relationship (cointegration) among selected variables. To do this, we performed the Pedroni [42,43] and Westerlund [44] panel cointegration tests. These tests are recommended when inter-country convergence is confirmed [45]. The final step of the empirical analysis is estimating the long-run coefficients (elasticities) of the cointegration association concerning urbanization.

For that purpose, we employed the panel Fully Modified Ordinary Least Squares test, the Pesaran and Smith [46] Mean Group estimator, the Pesaran [47] Common Correlated Effects Mean Group, and Augmented Mean Group estimators.

Step 1	Cross-Sectional Dependence Test	Pesaran CD test (2004)
Step 2	Panel Unit Root Test	Im, Pesaran and Shin test (2003) Pesaran (2007)
Step 3	Panel Cointegration Tests	Pedroni (1999, 2004), Westerlund (2005) tests
Step 4	Long-Run Coefficients Estimation	FMOLS, MG, CCEMG, AMG

**Figure 1.** The model estimation method. Notes: FMOLS—Fully Modified Ordinary Least Squares test; MG—Pesaran and Smith Mean Group estimator; CCEMG—Pesaran Common Correlated Effects Mean Group estimator; AMG—Augmented Mean Group estimator.

Our study sample consists of 28 countries for which we have complete data for the 2000–2018 period (532 observations for each variable). All data were retrieved from the World Bank and Eurostat databases. Table 1 describes the variables and sources of data. Table 2 shows the summary statistics. It should be noted that the differences between the values of the variables are appreciable in our research sample. The CO<sub>2</sub> emissions range between 2.97 metric tons per capita in Latvia and 25.67 in Luxemburg, while GDP per capita ranges between USD 3668.65 in Bulgaria and 105,454.7 in Luxemburg. At the same time, we observe considerable differentiation in the urban population, where the smallest values occur in Slovenia (59.7), and the biggest in Belgium (98.0).

**Table 1.** Variables descriptions and sources of data.

Variable	Description	Data Source
	dependent variable	
CO <sub>2</sub>	carbon dioxide emissions (metric tons per capita)	WDI
GHG	greenhouse gas emissions per capita. Indicator expressed in units of CO <sub>2</sub> equivalents in metric tons per capita	EEA
	Independent variable	
GDP	gross domestic product per capita (constant 2015 USD)	WDI
URB	urban population (% of total population)	WDI
ENC	tonnes of oil equivalent per capita	WDI

Notes: WDI—World Development Indicators; EEA—European Environment Agency.

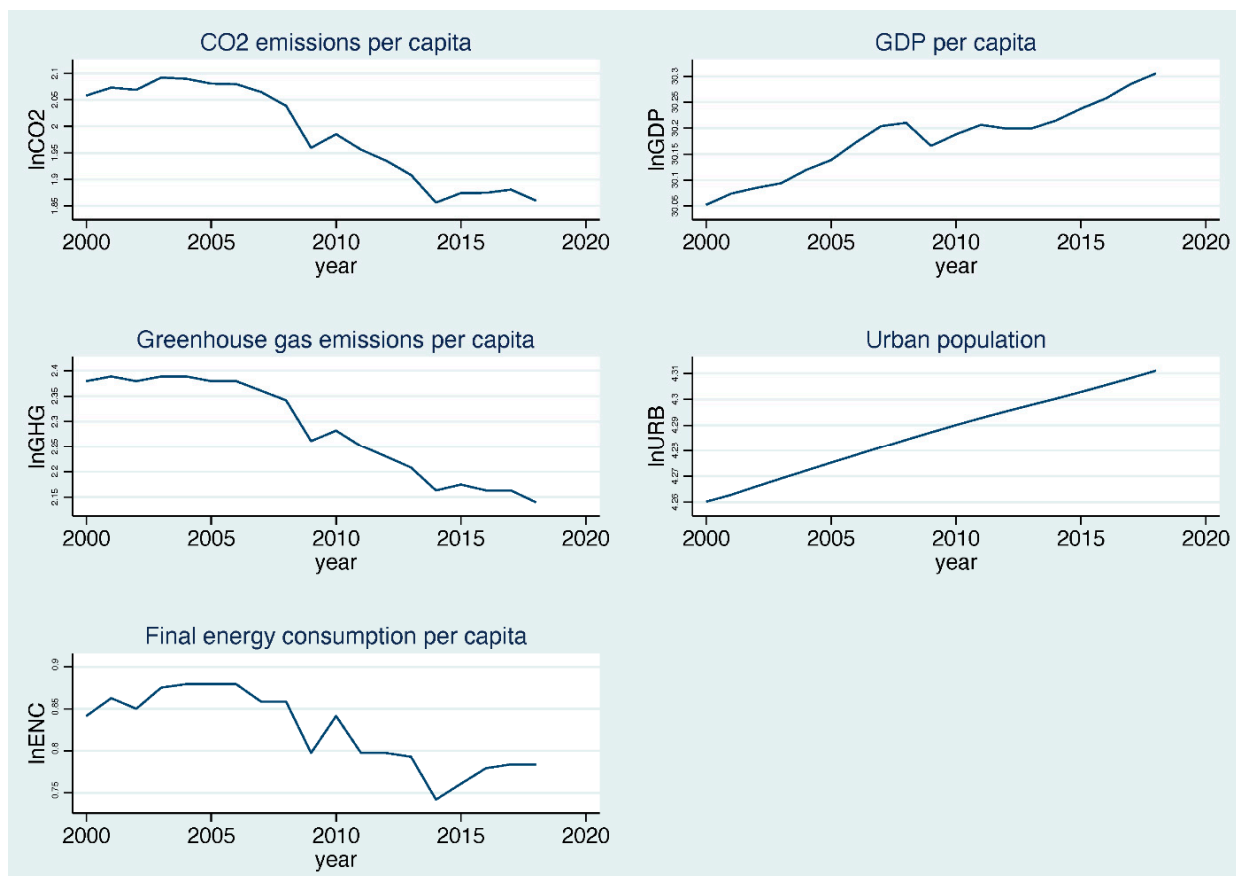
**Table 2.** Summary statistics.

Variable	Mean	Std. Dev.	Min.	Max.	Variance	Kurt.
CO <sub>2</sub>	8.104	3.975	2.927	25.669	15.799	6.260
GHG	10.426	4.157	4.3	30.8	17.278	8.304
GDP	29,163.460	19,691.830	3668.654	105,454.7	3.880	6.012
URB	72.140	12.444	50.754	98.001	154.861	2.129
ENC	2.487	1.372	0.930	9.630	1.883	12.648

Notes: CO<sub>2</sub>—carbon dioxide emissions (metric tons per capita); GDP—GDP per capita (constant 2015 USD); URB—Urban population (% of total population); ENC—Energy final consumption (tonnes of oil equivalent) per capita.

Figure 2 presents changes in aggregated variables for the European Union countries in the years 2000–2018. In the examined period, a decrease in per capita CO<sub>2</sub> and greenhouse

gas emissions, final energy consumption, and an increase in GDP per capita and urban population occurred. These trends should be assessed in a positive way. Another issue is a growing proportion of the population living in cities (urbanization), which is undoubtedly connected with the demographic changes, economic development, and technological advance observed in economically developed countries.



**Figure 2.** CO<sub>2</sub> emissions per capita, greenhouse gas emissions per capita, GDP per capita, urban population, and final energy consumption per capita in the European Union between 2000 and 2018.

#### 4. Results and Discussion

As outlined above, the first step in our method is to observe whether series are generating common shocks in the long run. To this aim, we test for cross-section dependence in our panel time-series data. The outcomes of the Pesaran cross-sectional dependency test [39] are shown in Table 3. The test results rejected the null hypothesis and confirmed the presence of cross-country dependency, which is not unexpected because the European Union countries share a common market and economic policy. A number of the conducted studies point to systematic economic convergence between these countries in recent years, for example, see Jóźwik [48] or Bernardelli et al. [49]. Because of this convergence, one country's economic and environmental transformations can easily be transferred to its neighboring countries. Therefore, we need to use a proper stationarity approach to circumvent the common effect and provide reliable results [45].

In the second step, we identify the order of integration of the variables by employing the Im–Pesaran–Shin panel unit root test. We subtracted the cross-sectional averages from the series and requested that the number of lags of the series be chosen in such a way that the AIC for the regression is minimized (max AIC is four). The stationarity test results in Table 4 confirm that the data series is unstable at this level. However, after considering the first difference, the test confirmed that the series became stationary at the 1% significance.

**Table 3.** Results of cross-sectional dependency Pesaran test.

Variable	Cd-Test	p-Value	Corr	Abs (Corr)
ln CO <sub>2</sub>	45.12 ***	0.000	0.532	0.666
lnGHG	40.12 ***	0.000	0.473	0.711
lnGDP	58.27 ***	0.000	0.688	0.768
lnGDPsq	58.26 ***	0.000	0.687	0.768
lnURB	26.65 ***	0.000	0.314	0.843
lnENC	24.48 ***	0.000	0.289	0.565

Note: Under the null hypothesis of cross-section independence  $CD \sim n(0,1)$ . \*\*\* denotes statistical significance at the 1% level.

**Table 4.** Im–Pesaran–Shin Panel unit root test (W-t-bar statistics).

Variables	Drift and No Trend	p-Value	Drift and Trend	p-Value
at level				
ln CO <sub>2</sub>	−0.8740	0.1910	−3.8365 ***	0.0001
lnGHG	−0.6831	0.2473	−4.6078 ***	0.0000
lnGDP	−0.0958	0.4618	−1.2311	0.1091
lnGDPsq	−0.2414	0.4046	−1.4185 **	0.0780
lnURB	−0.7913	0.2144	4.1977	1.0000
lnENC	−0.7915	0.2143	−4.6526 ***	0.0000
at 1st difference				
ln CO <sub>2</sub>	−15.6464 ***	0.0000	−10.9659 ***	0.0000
lnGHG	−14.9536 ***	0.0000	−10.1475 ***	0.0000
lnGDP	−6.6989 ***	0.0000	−4.1460 ***	0.0000
lnGDPsq	−6.7379 ***	0.0000	−4.5406 ***	0.0000
lnURB	−2.4104 ***	0.0080	−6.2014 ***	0.0000
lnENC	−14.1526 ***	0.0000	−8.6068 ***	0.0000

Notes: H0: All panels contain unit roots. Ha: Some panels are stationary. Cross-sectional means removed. Max AIC is 4. The number of lags of the series is chosen in such a way that the AIC for the regression is minimized. \*\*, \*\*\* denote statistical significance at the 5% and 1% levels, respectively.

In addition, we employed the panel second generation unit root test in the presence of cross-section dependence proposed by Pesaran [41]. We assume that the serial correlation order to be tasted with the Breusch–Godfrey Lagrange multiplier test in each regression is one, and the number of lags is four. Table 5 displays results for two deterministic models’ specifications: with individual-specific intercepts and incidental linear trends. The test results confirm that the variables are stationary at the first difference, almost all at the 1% significance.

**Table 5.** Pesaran panel unit root test in the presence of cross-section dependence.

Variables	At Level		At 1st Difference	
	Individual-Specific Intercepts	Incidental Linear Trends	Individual-Specific Intercepts	Incidental Linear Trends
ln CO <sub>2</sub>	−2.065	−2.801 **	−4.291 ***	−4.334 ***
lnGHG	−2.202 **	−3.034 ***	−4.331 ***	−4.460 ***
lnGDP	−2.023 **	−2.183 **	−3.024 ***	−3.068 ***
lnGDPsq	−1.978	−2.160 **	−2.941 ***	−3.060 ***
lnURB	−0.903	−1.126	−2.174 **	−3.680 ***
lnENC	−1.436	−2.908 ***	−4.205 ***	−4.171 ***

Notes: critical values: at \*\*\*—1% significant level is −2.32 and at \*\*—5% is −2.15; the serial correlation order to be tasted with the Breusch–Godfrey Lagrange multiplier test in each individual regression is 1; the number of lags is 4.

As noted previously, environmental degradation can be proxied in various ways. We selected CO<sub>2</sub> emissions as a proxy for environmental degradation in the model I, but for a robustness check, we also used the greenhouse gas emissions per capita variable. In this respect, we checked the cointegration (long-run relationship) between variables using the Pedroni and Westerlund tests. The tests have a common null hypothesis of no cointegration.



The alternative hypothesis of the Pedroni tests is that the variables are cointegrated in all panels. In the version of the Westerlund test in which the AR parameter is panel specific, the alternative hypothesis is that the variables are cointegrated in some of the panels. In the version of the Westerlund test in which the AR parameter is the same over the panels, the alternative hypothesis is that the variables are cointegrated in all the panels. In the Pedroni tests, we subtracted the cross-sectional averages from the series and requested that the number of lags of the series be chosen in such a way that the AIC for the regression is minimized (max AIC is four), as in the panel unit root tests calculations. Table 6 reports results for cointegration where six out of seven Pedroni tests confirm cointegration in Model I and Model II. The results of the Westerlund tests indicate that the variables are cointegrated in some of the panels.

**Table 6.** Pedroni and Westerlund panel cointegration tests results.

Tests	Model I	Model II
Pedroni test AR parameter: Same		
Modified variance ratio	−3.2956 ***	−3.5891 ***
Modified Phillips–Perron t	0.9245	1.1932
Phillips–Perron t	−8.0501 ***	−6.8419 ***
Augmented Dickey–Fuller t	−10.4652 ***	−10.1033 ***
Pedroni test AR parameter: Panel specific		
Modified Phillips–Perron t	3.0399 ***	3.1541 ***
Phillips–Perron t	−8.6913 ***	−7.3555 ***
Augmented Dickey–Fuller t	−12.4472 ***	−11.0408 ***
Westerlund test AR parameter: Same		
Variance ratio	−1.2798	−1.3705 *
Westerlund test AR parameter: Panel specific		
Variance ratio	−2.5575 ***	−2.6808 ***

Notes: Westerlund test AR parameter: Same. Ha: All panels are cointegrated; Panel specific. Ha: Some panels are cointegrated. \*, \*\*\* denote statistical significance at the 10% and 1% levels, respectively.

In the final step, we estimated the coefficients of Equations (1) and (2). Table 7 provides the FMOLS test results. To test the robustness of the estimated results, we used the Pesaran and Smith Mean Group estimator, the Pesaran Common Correlated Effects Mean Group, and the Augmented Mean Group estimators. These tests, which concern with correlation across panel members (cross-section dependence), were introduced by Eberhardt and Teal [50] and Bond and Eberhardt [51]. The advantage of using these estimators is that they are designed for ‘moderate-T and moderate-N’ macro panels. These results are presented in Table 8.

**Table 7.** Panel FMOLS test results.

Variable	Coefficient	t-Stat	p-Value
model I			
lnGDP	29.81 ***	813.83	$p < 0.00001$
lnGDPsq	−1.41 ***	−854.28	$p < 0.00001$
lnURB	−5.02 ***	−123.71	$p < 0.00001$
lnENC	1.05 ***	248.03	$p < 0.00001$
model II			
lnGDP	17.65 ***	702.78	$p < 0.00001$
lnGDPsq	−0.82 ***	−752.76	$p < 0.00001$
lnURB	−6.54 ***	−144.88	$p < 0.00001$
lnENC	0.83 ***	209.09	$p < 0.00001$

Notes: The number of observations is 532.  $p$ -value for two-tailed hypothesis. \*\*\* denotes statistical significance at the 1% level.

**Table 8.** Mean Group (MG), Augmented Mean Group (AMG), and Common Correlated Effects Mean Group (CCEMG) estimation results.

Test	Coefficient				
	lnGDP	lnGDPSq	lnURB	lnENC	Const.
	model I				
MG	44.505 ** (0.023)	−2.038 ** (0.024)	−4.289 *** (0.002)	0.994 *** (0.000)	−223.513 ** (0.038)
AMG	25.251 * (0.076)	−1.176 * (0.083)	−3.688 (0.205)	0.907 *** (0.000)	−118.128 (0.121)
CCEMG	−10.037 (0.350)	0.511 (0.302)	−6.895 (0.159)	0.951 *** (0.000)	57.116 (0.330)
	model II				
MG	28.605 ** (0.041)	−1.290 ** (0.047)	−6.067 *** (0.004)	0.8170 *** (0.000)	−130.276 * (0.087)
AMG	29.452 ** (0.019)	−1.375 ** (0.021)	−10.449 (0.194)	0.672 *** (0.000)	−109.249 (0.076)
CCEMG	−10.037 (0.350)	0.511 (0.302)	−6.895 (0.159)	0.951 *** (0.000)	57.116 (0.330)

Notes: numbers in parentheses are  $p$ -value. \*\*\*, \*\*, \* denote statistical significance at the 1%, 5%, and 10% level, respectively.

As can be seen from Tables 7 and 8, in both models, the significant coefficients of the real GDP per capita are positive, whereas those of the squared GDP per capita are negative. It means that the long-run linkage between CO<sub>2</sub> emissions per capita and GDP per capita is an inverted U-shape implying that the environmental Kuznets curve concept is verified for the whole group of the European Union countries. The economic growth and development are supportive of carbon emissions in that region. Similar results of studies for the group of European countries were recently obtained by Destek et al. [20], Maneejuk et al. [9], and Verbič et al. [19], as well as by Balsalobre-Lorente et al. [18]. Despite significant technological advances in the European Union countries, energy consumption still positively influences carbon emissions per capita. Our results are similar to the papers mentioned earlier, as well as to those that have been published recently, namely, Khezri et al. [17], Kar [52], and Mohsin et al. [53].

Interestingly, all results from Tables 7 and 8 indicate that urbanization negatively impacts carbon emissions per capita. However, there are differences in the significance level of the coefficients depending on the method used. Nevertheless, this shows that urbanization has an essential effect on environmental protection in the European Union area, nowadays. For example, the results of FMOLS for model I show that a 1% increase in a share of the urban population decreases emissions per capita by 5.02% if all other variables remain the same. We want to highlight that our study on the urbanization process' effect on carbon dioxide emissions points to different results than many studies mentioned in the literature review section. As we remember, the significant results indicate that urbanization positively impacted the carbon dioxide emissions in different groups of European countries. To give an example, in an article by Ali et al. [15], coefficients are positive and equal to 0.188 and 0.011; in the study by Balsalobre-Lorente et al. [18], between 0.44 and 6.36; while in the studies by Destek et al. [20] and Amin et al. [21] there was no relationship. This difference probably results from a few reasons, some of which include research methods, samples, and periods. Another reason can be related to trespassing the threshold after which both increased income per capita and the coefficient of urbanization give rise to improved quality of the environment, which was indicated by Gieraltowska et al. [11]. This effect can be enhanced by the deindustrialization process we wrote about in the literature review. Dong et al. [54] highlighted that from the perspective of income level, industrialization contributes to the growth in carbon emissions. The effect of industrialization on CO<sub>2</sub> emissions gradually increases in the low- and intermediate-income

levels. Azam et al. [55] also state that the industrialization process in OPEC economies increases environmental pollution, while the impacts on income are the opposite. However, the effect of industrialization begins to weaken at the high-income level according to research conducted by Dong et al. [54]. Probably this effect is observed in the European Union countries with a high-income level. Furthermore, economic development supports human capital, which significantly improves environmental quality [56]. Thus, our results indicate that studies in this area should be extended to different research models and methods.

## 5. Conclusions and Recommendations

In our research, we took into consideration two trends. First is the urbanization process, which increased the urban population in most European Union countries in years 2000–2018, and the second trend is a decrease in carbon dioxide emissions, which is indirectly the consequence of technological advances and the applied European climate policy. Considering these two trends, our research goal was to answer the following question: is there a long-run relationship between urbanization, energy consumption, economic growth, and carbon dioxide emissions, and what roles do urbanization and energy consumption play in the concept of the environmental Kuznets curve in European Union countries? We used the data from 28 European Union countries to assess the relationships. Our findings confirmed the long-run relationship between variables. We validated the environmental Kuznets curve hypothesis, indicating that economic growth has an inverted U-shaped effect on CO<sub>2</sub> emissions.

However, energy consumption still positively influences carbon emissions per capita, even though European Union countries have made significant economic and technological progress. At the same time, urbanization has a highly negative impact on carbon dioxide emissions per capita. If all other variables remain the same, a 1% increase in a share of the urban population decreases CO<sub>2</sub> emissions per capita by 5.02%. The result of our study is different from the results in the majority of earlier published articles. This difference probably arises from a few reasons. One of them may be the fact that the threshold after which both an increase in income per capita and urban population causes a decrease in carbon dioxide emissions in European Union countries has been trespassed in recent years.

Our results provide new insights for policymakers in European Union institutions. The findings suggest that the European policy should support the process of urbanization in a complex manner to fulfill the European Green Deal and achieve the Sustainable Development Goals related to improving environmental quality, especially by promoting urbanization with a low-carbon infrastructure and transport (smart technology and energy-efficient hybrid vehicles). A positive coefficient associated with energy consumption indicates that local authorities should support the development of home renewable energy infrastructure, for example, energy-efficient electric appliances and solar energy. Another important practical implication is related to human capital. The urban population can be motivated to adopt a sustainable lifestyle, including energy-saving, renewable energy sources, and public transportation [56]. It is very important in this context that urbanization be carried out according to environmental norms, possibly without social compromises in this respect. In addition, modern technological solutions enable the development of intelligent cities that are environmentally neutral.

However, we only conducted a preliminary empirical analysis of the relationship between environmental degradation and urbanization, and our study has a few limitations. The first limitation refers to sample size. The sample covers the period 2000–2018, this means we should be cautious in generalizing the findings. Second, although we have robust results using an alternative measure of environmental degradation, the two proxies (CO<sub>2</sub> and greenhouse emissions) might limit the ecological degradation effects. Additionally, it would be interesting to examine the consumption environmental impacts offshored to other countries and the deindustrialization processes. Third, we did not divide the European Union countries, for example, into less developed countries (Central European countries)

and developed countries (Western European countries) to make a comparative analysis. These limitations could be addressed in future research.

Undoubtedly, in further research, we must also remember that climate neutrality is a global challenge. This means that it requires international dialogue and cooperation between the states. Although the pressure applied usually refers to particular countries and their economic structures, international activity is also an issue that plays a predominant role. It is especially important due to the necessity of creating a synergy between the European and international climate initiatives. For this reason, understanding that adaptation to climate changes is important; however, this is not in itself the aim, but rather a principle. It should, however, be a component of properly functioning and developing countries and societies.

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## Appendix A

**Table A1.** CO<sub>2</sub> emissions per capita and urban population in the European Union countries.

Country	CO <sub>2</sub> Emissions (Metric Tons per Capita)		Urban Population (% of Total Population)	
	2000	2018	2000	2018
Austria	17,690	15,476	60,213	58,297
Belgium	11,441	8180	97,129	98,001
Bulgaria	5303	5855	68,899	75,008
Croatia	4040	4056	53,428	56,947
Cyprus	7495	6079	68,648	66,810
Czech Rep.	12,011	9641	73,988	73,792
Denmark	12,011	9641	85,100	87,874
Estonia	10,609	12,103	69,368	68,880
Finland	10,645	8043	82,183	85,382
France	6127	4619	75,871	80,444
Germany	10,097	8558	74,965	77,312
Greece	8742	6083	72,716	79,058
Hungary	5350	4746	64,575	71,351
Ireland	11,201	7624	59,155	63,170
Italy	7662	5376	67,222	70,438
Latvia	2927	3959	68,067	68,142
Lithuania	3003	4137	66,986	67,679
Luxembourg	19,665	15,330	84,216	90,981
Malta	5460	3198	92,368	94,612
Netherlands	10,191	8773	76,795	91,490
Poland	7729	8235	61,716	60,058
Portugal	5992	4841	54,399	65,211
Romania	3960	3845	53,004	53,998
Slovak Rep.	7065	6059	56,233	53,726
Slovenia	7310	6775	50,754	54,541
Spain	7230	5520	76,262	80,321
Sweden	6005	3538	84,026	87,431
United Kingdom	9001	5399	78,651	83,398

Source: World Development Indicators.

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